

Locked Out of College: When Admissions Bureaucrats Do and Do Not Discriminate *

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Abstract

How does a criminal record shape interactions with the State and society? We present evidence from a nationwide field experiment in the United States, showing that prospective applicants with criminal records are about five percentage points less likely to receive information from college admission offices. However, we demonstrate that bias does not extend to race. There is no difference in response rates to Black and White applicants. We further show that bias is all but absent in public bureaucracies, as discrimination against formerly incarcerated applicants is driven by private schools. Examining why bias is stronger for private colleges, we demonstrate that the private-public difference persists even after accounting for college selectivity, socio-economic composition and school finances. Moving beyond the measurement of bias, we evaluate an intervention aimed at reducing discrimination: whether an email from an advocate mitigates bias associated with a criminal record. However, we find no evidence that advocate endorsements decrease bureaucratic bias.

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Does a criminal record inhibit access to social goods and services? Politics has long been understood as “who gets what, when, and how” (Lasswell, 1936), but social scientists have only recently begun to explore how punitive policies influence this power structure (Manza and Uggen, 2006; Weaver and Lerman, 2010). The State’s power to punish is augmented by its ability to attach stigmatic labels to legal transgressors (Foucault, 1977), labels that influence individuals’ interactions with the State and society post-incarceration. Directly or indirectly, a criminal record can inhibit access to voting rights, employment, welfare and education (Pager, 2003; Manza and Uggen, 2006; Owens and Smith, 2012). Even where access is not explicitly prohibited, formerly incarcerated persons often have to disclose their criminal record. These requirements complicate the bureaucratic process and open opportunities for discrimination.

To what extent, and in what context, do people with criminal records face discrimination? Experiments are frequently used in the social sciences to measure discrimination in access to social goods, including voting registration, public housing, employment, and medical services (Bertrand and Mullainathan, 2004; Pager, Bonikowski and Western, 2009; White, Nathan and Faller, 2015). Most of these studies focus on racial discrimination, generally finding that minorities face discrimination in accessing goods and services (but see Einstein and Glick (2017)). Only a few experiments examine discrimination based on criminal record, primarily in hiring (Pager, 2003). Discrimination in higher education – an important determinant of political participation and recidivism – is understudied. Formerly incarcerated populations face deficits in educational attainment, and Blacks and other minority groups are under-represented in higher education (Hjalmarsson, Holmlund and Lindquist, 2015). Punitive labeling by one state institution – the penal system – could negatively affect access to this social good that is also partially provided by the state.

We test for discrimination against formerly incarcerated college applicants through a randomized field experiment. We send emails to 2,917 college admissions offices, inquiring about the requirements for application and admission. We use a factorial design with three treatments, which we present in Table 1 and Figure B2. In each email, the applicant reveals that he¹ has a General

¹Given that the incarcerated population in the US is overwhelmingly male, the applicants are always men.

Education Diploma (GED) and asks about eligibility for admission. The emails are randomly assigned with equal probability to disclose that the applicant got their GED either online or in a state penitentiary. To test for racial bias, the applicant is randomly assigned to have a putatively White or Black name. Our primary outcome of interest is the rate of response for different treatment conditions. The overall response rate is 74%. Recognizing that bias can be multidimensional, we also consider two additional outcomes: friendliness and thoroughness of the response.

Table 1: Overview of Treatment Conditions

| Treatment | Possible values | Implementation |
|------------------|---|-----------------------------------|
| Criminal record | {Criminal record, no criminal record} | GED obtained in prison or online |
| Race | {Black, White} | Signaled via name of applicant |
| Advocate | {No advocate, Black advocate, White advocate} | Email is sent by a former teacher |

Note: This table describes the implementation of the three treatment conditions. The first column is the name of the treatment. The second column lists all possible values for each treatment. The third column summarize how the treatment was implemented.

Moving beyond measurement of bias, we propose an intervention to help mitigate discrimination. We test whether the support of an advocate, in the form of a former teacher, can help marginalized populations extract information from potentially biased bureaucracies. Assuming that former teachers are unlikely to vouch for unqualified candidates, their endorsement can serve as a signal of applicant quality. If bias against formerly incarcerated individuals occurs when admission bureaucrats use a criminal record to proxy for unobserved characteristics of the applicant, an advocate can serve as an added credential (Gaddis, 2014). To implement the advocate intervention, we randomize whether the email is sent by the applicant himself, or whether it is sent by a former GED instructor of the applicant. This is a low-cost intervention: Using a short and simple email, the teacher reaches out on behalf of the applicant to inquire about college eligibility. To increase comparability between the applicant and advocate emails, we also randomize the race of the advocate. A Black applicant can have either a Black or a White advocate, and vice versa. In Table 1 and Figure B2, we document the three treatment conditions.

Admissions officers have significant discretion over whether and how they reply to prospective applicants. While colleges have admissions policies governing how criminal record can or cannot

be considered, these policies do not explicitly extend to information provision. Furthermore, admissions policies themselves are often ambiguous, creating the opportunity for admissions officials to use their own discretion when interacting with potential applicants. Few colleges specifically bar admission of applicants with felony criminal records, but almost all colleges retain the right to refuse admission based on past criminal activity, and often require additional forms or essays of formerly incarcerated applicants². Thus, admissions eligibility is often unclear, and formerly incarcerated applicants must overcome this first hurdle before applying.

We stress that our design measures bias in bureaucratic responsiveness, rather than bias in admissions. Feasibly, admissions bureaucrats may have different incentives when responding to general inquiries versus when actually making the decisions to admit or reject an applicant³. Still, bias in bureaucratic responsiveness at this entry point into the admission process is likely to depress enrollment for formerly incarcerated applicants. Non-responsiveness inhibits an applicant's ability to get important information as to what is required to enroll. Given the additional bureaucratic procedures for applicants with criminal records, ascertaining eligibility is an important first step to enrollment. The relevance of information in college admissions is further underlined by a growing number of studies that examine interventions aimed at increasing access to admission information (Bettinger et al., 2012; Hoxby and Turner, 2015; Deming and Dynarski, 2010; Dynarski et al., 2018). This literature finds that a lack of information strongly decreases the probability of application, enrollment and eventual success in college, particularly for applicants from lower socioeconomic backgrounds. What is more, applications in which formerly incarcerated applicants struggle to access crucial information can cause such applicants to self-select out of the process (Rosenthal et al., 2015). Access to information is increasingly seen as a key determinant of success in higher education, highlighting the need to study bias in bureaucratic responsiveness to requests

²In August 2018, the Common Application removed the section asking about criminal record. This new policy took effect after the running of this experiment. Many colleges still require disclosure of criminal record in separate forms.

³A recent study by Stewart and Uggen (2020) tests for discrimination based on criminal record in admissions decisions, finding a strong negative effect.

for admission information.

Drawing on the results from our field experiment, we offer several contributions: This is the first study, to the best of our knowledge, to establish the causal effect of a criminal record on bureaucratic responsiveness in the context of higher education. On average, formerly incarcerated individuals are about 5 percentage points less likely to receive a response from admissions offices. However, further results paint a more positive picture: First, even though we send queries from putatively low-SES applicants, response rates are relatively high, at about 75%. Second, contrary to a large body of empirical work, we demonstrate that bias does not extend to applicants' racial backgrounds: There is no difference in response rates for Black and White applicants.

Exploring treatment effect heterogeneity, we demonstrate that institutional context is a key mediator of bias. Public institutions do not discriminate against prospective applicants with criminal records. Bias in response is driven by private colleges, which are about 10 percentage points less likely to respond than their public counterparts. In an additional analysis that was not pre-registered, we explore four different explanations for why bias is more prevalent in private institutions: (1) Differences in admissions selectivity, (2) the socio-economic makeup of the student bodies, (3) differences in school finances and (4) institutional priorities. We test the first three of them using additional school characteristics from Chetty et al. (2017), but find no evidence for either of these explanations. We do, however, find suggestive evidence that differences in treatment effects between public and private schools can be explained by private schools being far more likely to require disclosure of an applicant's criminal record. Admissions bureaucrats may internalize institutional priorities signaled by these policy differences, potentially explaining the differential treatment of formerly incarcerated applicants.

Our empirical results extend the literature in a number of ways: First, we demonstrate the mark of a criminal record for a good that is provided both publicly and privately: higher education. Second, we document that context matters, with admissions bureaucrats at private institutions exhibiting non-responsiveness towards applicants who have spent time in prison at a much greater rate than admissions bureaucrats at public universities. In addition, we move beyond measurement

of bias to test a strategy for reducing bias, but do not find evidence that an advocate reaching out on an incarcerated applicant's behalf is effective in this context.

1 Empirical Strategy

As laid out in the pre-analysis plan, we contact 2,917 public and private non-profit colleges across the United States. We constructed this sample from an exhaustive list of colleges operating in 2018, obtained from the National Center for Education Statistics (NCES). Each college admissions office receives one email, inquiring about information concerning the admissions process. The emails express interest in applying to the college, then report that the applicant has his GED, and ask if that affects his eligibility. Finally, the emails inquire what else is required to apply and whether the college is currently accepting applications.

Each email is randomly assigned to treatment or control groups across three conditions: *Criminal record*, *race* and *presence of an advocate*. The first two treatment conditions are binary. The applicant either has a criminal record or not, and the applicant is either Black or White. About 50% of all colleges receive emails that are sent directly by the applicant. The remaining emails are sent by a former teacher of the applicant (the advocate), who is either Black (25% of all emails) or White (25% of all emails). The former teacher inquires on behalf of the applicant without explicitly endorsing the applicant. The advocate treatment can take on three values: no advocate, Black advocate or White advocate. Both the presence of the advocate and the advocate race is independent of the applicant's race or criminal record. Our design results in $2 \times 2 \times 3 = 12$ different treatment arms. In Table 1, we present an overview of the treatment conditions, while Figure B2 in the supplementary information (SI) shows the email language across treatment combinations. In Table 2, we present average response rates and number of emails sent for each of the 12 possible treatment arms.

We use the location at which the GED was obtained to signal that the applicant has spent time in prison. The applicant either reports that he received his GED at a state penitentiary (treatment)

Table 2: Treatment Arms, Response Rates and Number of Emails

| Criminal Record Treatment | Race Treatment | Advocate Treatment | Response Rate | N |
|---------------------------|----------------|--------------------|---------------|-----|
| No Criminal Record | White | No Advocate | 0.751 | 360 |
| No Criminal Record | White | White Advocate | 0.856 | 174 |
| No Criminal Record | White | Black Advocate | 0.783 | 190 |
| No Criminal Record | Black | No Advocate | 0.810 | 397 |
| No Criminal Record | Black | White Advocate | 0.808 | 172 |
| No Criminal Record | Black | Black Advocate | 0.803 | 174 |
| Criminal Record | White | No Advocate | 0.748 | 346 |
| Criminal Record | White | White Advocate | 0.756 | 203 |
| Criminal Record | White | Black Advocate | 0.746 | 194 |
| Criminal Record | Black | No Advocate | 0.737 | 364 |
| Criminal Record | Black | White Advocate | 0.747 | 175 |
| Criminal Record | Black | Black Advocate | 0.776 | 185 |

Note: This table reports mean response rates and number of observations for all 12 possible treatment combinations.

or online (control)⁴. Signaling that the applicant has a GED allows us to keep socio-economic status constant across treatment conditions, thus lowering the probability that admissions officers conflate the treatments with social class. At the same time, it is feasible that having completed a GED in prison signals serious commitment to education. Admission officers may ascribe greater motivation or commitment to applicants who completed their GED in prisons. Therefore, our treatment could result in *higher* response rates for applicants who completed their GED in prison.

Moving to the race treatment, we use either a putatively Black (Tyrone Booker, Darnell Banks) or White (Kevin Schmidt, Bob Krueger) name to reveal race. We chose names that were not used in previous audit studies, and pre-tested those names for consistent racial connotations (we present the pre-test results in Table B1 in the SI).

The advocate treatment changes the email language slightly: Instead of the applicant, a named advocate identifies himself as a former teacher of the applicant and then proceeds to inquire about the exact same information as in the standard direct email. The advocate is either White or Black

⁴The state penitentiaries used in the emails were minimum security prisons in the same state as the college. We include state fixed effects in our empirical strategy to eliminate concerns that the effects are confounded by our choice of the penitentiary in each state.

with equal probability. As shown in Table 2, the advocate treatment is assigned equally across applicant race or criminal record. To rule out spurious effects that result from differences in email language, we have tried to make the applicant and advocate emails as similar as possible. Consequently, the advocate does not explicitly endorse the applicant beyond the fact that the advocate is inquiring on behalf of the applicant.

Using an automated script, we sent 2,934 emails over the course of eight weekdays between February 23, 2018 and March 6, 2018. Each day, we sent about 360 emails. We randomized the order in which emails were sent. 72 emails could not be delivered, since the email addresses that we obtained were not up to date, or, in a few cases, the college did not exist anymore. For this subset of emails, we collected new contact information, and tried resending all emails on March 12, 2018. For all except 17 colleges, this was successful, bringing our final sample size to 2,917.

Our main measure of bias is the responsiveness of the admissions offices, a binary variable coded as 1 if we received a response within 21 days and 0 otherwise. Recognizing that discrimination can be multidimensional (Hemker and Rink, 2017), we include two additional outcomes, thoroughness and friendliness. As in Einstein and Glick (2017), we conceptualize friendliness as a binary variable that we code as 1 if a response addresses the sender by name. Thoroughness is coded on a numeric scale from 0 to 3, based on whether the response answered the 3 questions posed in the email⁵. To avoid conditioning on a post-treatment variable (Coppock, 2018), non-responses for the friendliness and thoroughness outcomes are coded as 0.

To construct the sample, we collected contextual data from the Integrated Postsecondary Education Data System, administered by the NCES. The database contains information on whether each school is a public or private school, whether it is primarily a four-year or two-year institution, and the size of the student body. These definitions and classifications are established by the NCES. A breakdown of the covariate distributions is shown in Figure B1 in the SI. Four-year colleges constitute about two thirds of the full sample, there are slightly more public than private institutions,

⁵Thoroughness was coded by Mechanical Turk workers. Each response was coded independently by two workers. We use the average of the resulting ratings as the thoroughness outcome for each response.

and most schools have fewer than 5,000 students.

Each of these contextual variables could be associated with bureaucratic capacity, school policies or student recruitment strategies, which in turn may influence response rates. We use coarsened exact matching (CEM, see Iacus, King and Porro 2012 and Section A.2 in the SI) to ensure balanced treatment assignment across the three pre-treatment variables. Since the advocate race treatment depends on the presence of an advocate, we use simple randomization instead of pair matching to assign that treatment status. In Figure A3 in the SI, we show that balance across the covariates is almost perfect.

For each of the three outcomes, we run two baseline models: with and without covariates. We include controls for the three pre-treatment covariates used to define the CEM strata. Subsequently, we interact the treatments to test whether bias varies by race or by presence of an advocate. We also include non-parametric estimations of our main treatment effects, to demonstrate that the results are not model-dependent (Table B3 in the SI). We then test whether the treatment effects vary between public and private colleges. Finally, we conduct a number of unregistered additional analyses to explore why the effect of criminal records varies between public and private institutions in Section 3.

1.1 Ethical Considerations

In designing the experiment, we took several steps to address important ethical concerns about the burden that audit studies impose on bureaucratic institutions, as well as the potential impact on the populations that rely on these bureaucracies. To minimize administrative burden, we tailored the email language to be brief and with questions that do not require lengthy answers. The “does getting a GED while incarcerated affect eligibility” query and the “are you currently accepting applications” portion commonly elicited very brief affirmative responses, while the “what else is needed to apply” query was often answered by pointing to a website. In addition, we never corresponded with the admissions offices beyond the initial email.

Contacting a smaller number of colleges would have reduced the burden imposed by the study.

While the resulting sample size might have been sufficient to detect the main effects, our pre-registered interaction effects require a larger number of observations. Since (1) these interactions are substantively relevant (e.g. the interaction between applicant race and criminal record) and (2) since the individual burden for each school is low, we decided to utilize a relatively large sample.⁶

Our design utilizes deception, which risks alienating the study sample, potentially influencing future bureaucratic behavior. We rely on deception in our design because it is the only way to test for real-world bias in responsiveness, and we argue that our research questions are of significant social importance so as to warrant such design choices (see also Einstein and Glick, 2017, for a related discussion in a similar setting). Finally, to ensure anonymity, all of our analysis reports results in the aggregate, and no identifiable information, for schools or individuals, is reported or shared.

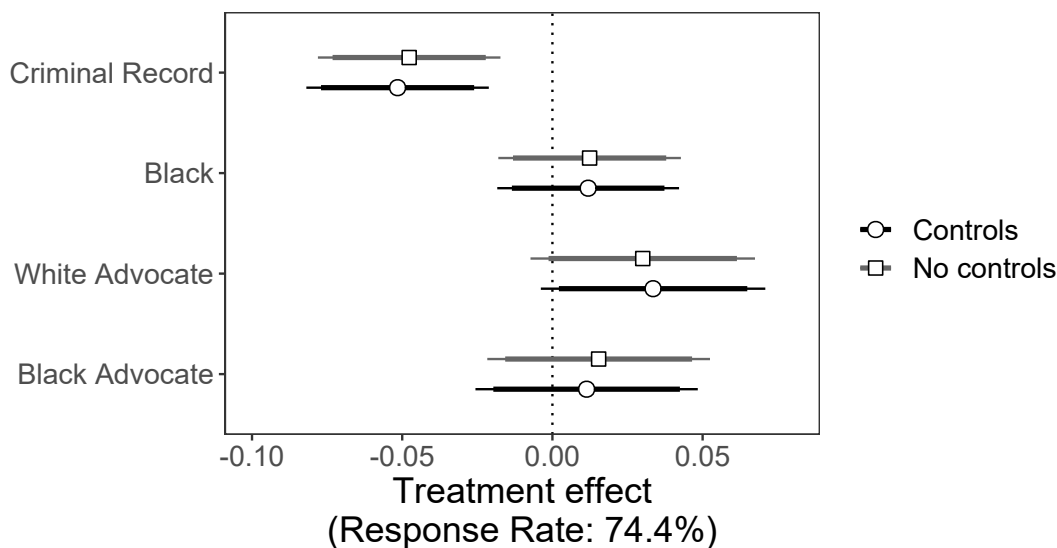
2 Results

In Figure 1, we present the main results. All else equal, admissions bureaucrats are 5.2 percentage points less likely to respond to a formerly incarcerated applicant. For applicants with putatively Black names, we do not find any evidence of bias in response rates. In fact, average response rates for Black applicants are somewhat higher than for White applicants, although these estimates are statistically indistinguishable from zero. In Table B7 in the SI, we present specifications that include interactions between the treatments. We do not have sufficient power to reject the null hypothesis for any of the interactions. However, we note that bias in responsiveness for formerly incarcerated applicants is stronger when the applicants are Black.

Regarding the advocate treatment, there is little evidence that the intervention increases response rates. While we observe increased response rates for White (but not Black) advocates in one specification, the interaction models (see Table B7 in the SI) show that the advocate effect decreases for applicants with criminal records.

⁶As we discuss in the results section, the observed interaction effects in Table B7 are often sizable, but not precisely estimated, even though the sample is already relatively large.

Figure 1: Main Results

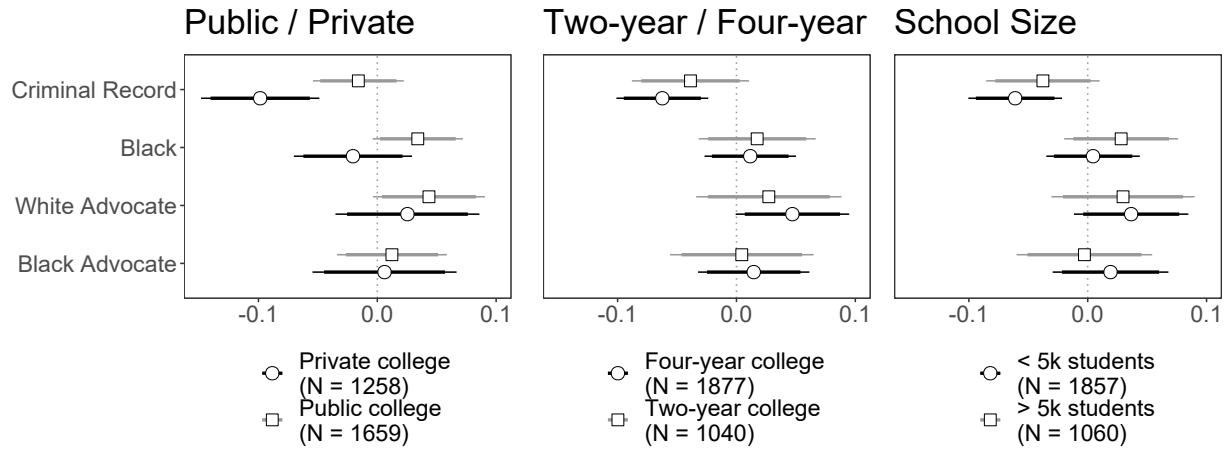


Note: The figure shows coefficient estimates from the main specifications. Each pair of coefficients refers to a treatment, which is shown on the y-axis. The outcome is a binary response indicator. Positive effect sizes indicate that the treatment condition increases response rates. The average response rate is 74.4%. The covariates are public/private, two-year/four-year, institution size and state fixed effects. The solid horizontal lines indicate 95% confidence intervals.

Having established that formerly incarcerated applicants are subject to bias, we examine whether the treatment effect varies with institutional characteristics in Figure 2. We observe the most pronounced heterogeneity when comparing bias at private versus public colleges. Private college admissions bureaucrats are about 10 percentage points less likely to reply to formerly incarcerated applicants. For public schools, there is no detectable difference in response rates. These estimates are statistically distinct from one another, as we show in Table B6 in the SI. The aggregate effects reported in Figure 1 are therefore driven by bias in private college admissions offices. We also find that private colleges do not appear to discriminate based on race, while public colleges tend to be more responsive to Black applicants. These effects are, however, not precisely estimated (significant only at $\alpha = 0.1$).

Since bias can be multidimensional, we also examine two alternative measures of bias: Thoroughness and friendliness. In Sections B.1 and B.2, we re-estimate all specifications discussed previously using these two outcomes. By and large, we observe the same patterns for the two alternative outcomes.

Figure 2: Results Conditional on Institutional Characteristics



Note: The figures show coefficient estimates, subset by school characteristics. The treatments are shown on the y-axis. The outcome is a binary response indicator. The average response rate is 74.4%. All specifications include covariates and state fixed effects. The covariates are two-year/four-year and institution size. The solid horizontal lines indicate 95% confidence intervals.

Given the large number of tests that we conduct, we adjust p-values using the Benjamini and Hochberg (1995) method to control the expected proportion of incorrectly rejected null hypotheses. We present adjusted p-values in Tables B19 and B20 in the SI. After adjusting for multiple comparisons, the observed negative effect of criminal records on response rates (see Figure 1) remains significant at $\alpha = 0.05$. The adjusted p-value for the interaction between criminal records and public schools increases to 0.063.

3 Differences Between Private and Public Institutions

What explains the disparity between public and private universities? While this statistical comparison was pre-registered, it was not one of the pre-stated primary focuses of the study. In response to the observed difference between response rates at public and private universities, we evaluate potential explanations for this difference, and test each explanation through additional unregistered analyses.

There are often tradeoffs in efficiency and equity between public and private provision of goods and services (Niskanen, 1968). For example, Jilke, Van Dooren and Rys (2018) conduct an audit study to examine ethnic bias in responsiveness among public and private Flemish nursing

homes, finding that discrimination against applicants with Maghrebian names is markedly more pronounced among private nursing homes. Additionally, an audit study of professor correspondence found greater unresponsiveness to minority and women students, particularly at private universities (Milkman, Akinola and Chugh, 2015). Regarding our findings, we propose four reasons for why the effect of a criminal record varies across public and private colleges:

Selectivity: Since prior incarceration is often associated with lower academic achievement (Blomberg et al., 2011), admission bureaucrats at more selective private colleges⁷ may not respond to inquiries if they believe that the probability of admitting the prospective applicant is low.

Socio-economic composition: Contact theory (Pettigrew, 1998) and familiarity bias (Tversky and Kahneman, 1974) suggest that bureaucrats may be more willing to assist applicants that they are familiar with. Public colleges may therefore be more responsive, as their student body consists of a greater number of students from racially and economically diverse backgrounds.

Financial considerations: Admissions bureaucrats may anticipate that formerly incarcerated or Black applicants are unable to pay for college without significant financial aid, so private schools, which are often more expensive and tuition-dependent, may exhibit lower levels of responsiveness.

Institutional priorities: There may be differences in overarching priorities of public versus private colleges to pursue diversity or support under-privileged groups. Public higher education policy in the US is founded on the principle of “[e]quality of opportunity for all students to attend public higher education in their state, without regard to their background or preparation” (Bastedo and Gumport, 2003, p. 341). In contrast, such a prerogative does not explicitly apply to private colleges, which might have different priorities when selecting prospective students. These differences in institutional priorities may be reflected in explicit policies, such as requiring applicant disclosure of a criminal record, that inform admission bureaucrat behavior.

To test the first three of these explanations, we merge our data with college data from Chetty

⁷In our sample, private colleges are, on average, about 4.2 percentage points more likely to reject applicants than public colleges.

et al. (2017). We test whether treatment effects are driven by admissions rejection rate or average SAT score (selectivity), racial demographics, parental median income, and percentage of parents in the top 1% of the nationwide income distribution (socio-economic composition), and tuition sticker price and net costs after financial aid (financial considerations). We find little evidence that selectivity, socio-economic differences or financial considerations underlie the public-private differences. In Tables B10 and B11 in the SI, we present the effect of time in prison on responsiveness subset by public and private, with interactions between the treatment the aforementioned variables from the Chetty et al. (2017) data. The magnitude of the treatment effect remains relatively unchanged despite the inclusion of interactions, and the coefficients on the interactions are all effectively zero. Thus, none of these characteristics appear responsible for the public-private differences in responsiveness.

To test whether differences in admissions policies reflect differences in admissions officer behavior, we merged our data with data from Stewart and Uggen (2020) on whether colleges require in their application disclosure of a criminal record. These data consist of 1,330 4-year universities and the policies are current as of 2015, while our experiment was run in early 2018. In this sample, private schools are far more likely to require disclosure of the applicant criminal record than public schools (Figure B3), and the treatment effect of a criminal record on response rate is largest (most negative) for schools that require disclosure (Table B12), although this discrepancy is imprecisely estimated. While not conclusive, this illustrates the possibility that private schools may be exhibiting bias against applicants with criminal records because of bureaucratic policies that prioritize this bias-inducing information.

4 Discussion

Using a nationwide randomized field experiment, we establish the causal effect of a criminal record on bureaucratic discrimination in college responsiveness. Punitive labeling by the state has demonstrable effects on access to higher education, a vital social good with downstream effects on political participation, labor market outcomes, and recidivism. While not a direct test for bias in

college admission, our results speak to a growing literature that highlights lack of information as a barrier to successful college admissions (see e.g. Hoxby, 2009), especially for low-SES applicants (Dynarski et al., 2018).

While we document bias against formerly incarcerated applicants, we also highlight two positive results: A high overall response rate and no evidence for racial bias in responsiveness. Even though we send queries from applicants with putatively low-SES, about three quarters of all inquires receive replies within three weeks. Unlike audit studies in other contexts (Costa, 2017), we do not find evidence of racial discrimination. We offer two possible explanations for this: First, all applicants have GEDs, which holds social class constant. If bias against putative Black applicants occurs when admissions bureaucrats conflate race with social class or education, holding education constant should eliminate some of this bias. This finding is consistent with Einstein and Glick (2017), who find no racial bias when applying to public housing – a setting where applicants are of similar socioeconomic strata and where minorities are disproportionately represented.⁸ Second, colleges may have successfully implemented policies to curb racial bias in admissions offices, e.g. because they strive for a diverse student body. Yet, while colleges may have been successful at curbing one dimension of bias in responsiveness, i.e. racial bias, more effort needs to be directed at decreasing bias against formerly incarcerated applicants.⁹

Akin to prior research on bureaucratic bias (Jilke, Van Dooren and Rys, 2018), our work highlights stark differences in the behavior of public versus private institutions. We find that discrimination by public colleges is close to zero, while private colleges discriminate at significantly higher rates. We propose four explanations for the observed public-private divide: (1) admissions selectivity, (2) the economic and racial makeup of the student body, (3) financial aid and dependence on

⁸This result is similar to another recent audit study of college admissions offices by Druckman and Shafranek (2020) testing for bias by partisanship and race. Their study also finds that applicant race does not affect response rates.

⁹In the context of our study, the fact that applicants completed their GEDs in prison may signal greater commitment or motivation to admission bureaucrats. In the absence of such a signal, bias against formerly incarcerated applicants could be even more pronounced.

tuition, and (4) differences in institutional priorities. To examine those explanations, we conduct several exploratory analyses that were not pre-registered. Using college characteristics compiled by Chetty et al. (2017), we find little evidence for the first three mechanisms. Using data from Stewart and Uggen (2020), we do find some evidence that colleges that require disclosure of applicant criminal records are more likely to exhibit bias against applicants with criminal records, and private schools are far more likely to require this disclosure than public schools. This suggests that admissions bureaucrats operate under general differences in bureaucratic priorities, and the fact that private schools are more likely to require disclosure of a criminal record signals to the admissions officers that this information may warrant differential treatment. Future research could further connect admissions policies to admissions bureaucrats' behavior, and better investigate why these differences in admissions policies emerge.

Moving beyond measurement, we propose a strategy for marginalized populations to navigate biased bureaucracies (see Butler and Crabtree, 2017). Instead of a direct inquiry by applicant, some emails are sent by advocates. However, we find little evidence that emails from former teachers of students can ameliorate the effects of a criminal record. While we find some evidence for increased response rates for White (but not Black) advocates, the interaction models (see Table B7 in the SI) show that the advocate effect is diminished greatly for applicants with criminal records. We offer two explanations for why the advocates were unable to reduce bias. First, the advocate email does not contain a direct endorsement of the applicant. Since our aim was to make emails comparable across advocates and applicants, we did not include an explicit endorsement. The advocate email might therefore be a weak signal of applicant quality – too weak to convince admissions bureaucrats. Second, applicants will likely have to submit teacher recommendations when they apply to college. Bureaucrats may expect that any potential applicant will be able to obtain endorsements from their teachers anyway, so the advocate email might not reveal additional information to bureaucrats. Although we have little evidence in favor of advocate endorsements as a strategy to reduce bias, we stress that future research should examine more nuanced implementations of the advocate intervention.

Our findings show that the absence of racial differences in responsiveness does not imply unbiasedness along other dimensions, such as applicants' criminal histories. Echoing the results in Einstein and Glick (2017), we find that bias may vary substantially, even when we hold constant the social good that is provided. Future researchers could examine why the same bureaucracy is biased along one dimension, but not the other. There are also other important types of biases that may exist and are worth identifying. For example, a recent study of secondary schools in Spain found bias in responsiveness against homosexual parents (Diaz-Serrano and Meix-Llop, 2016). In addition, we argue that future work on bureaucratic bias needs to emphasize institutional differences in service provision, as many goods and services are routinely provided both by private and public actors, such as housing, health care, transportation and education.

References

- Bastedo, Michael and Patricia Gumport. 2003. "Access to what? Mission differentiation and academic stratification in US public higher education." *Higher Education* 46(3):341–359.
- Benjamini, Yoav and Yosef Hochberg. 1995. "Controlling the false discovery rate: a practical and powerful approach to multiple testing." *Journal of the Royal Statistical Society: Series B (Methodological)* 57(1):289–300.
- Bertrand, Marianne and Sendhil Mullainathan. 2004. "Are Emily and Greg More Employable Than Lakisha and Jamal? A Field Experiment on Labor Market Discrimination." *American Economic Review* 94(4):991–1013.
- Bettinger, Eric P., Bridget Terry Long, Philip Oreopoulos and Lisa Sanbonmatsu. 2012. "The Role of Application Assistance and Information in College Decisions: Results from the H&R Block Fafsa Experiment*." *The Quarterly Journal of Economics* 127(3):1205–1242.
- Blomberg, Thomas G, William D Bales, Karen Mann, Alex R Piquero and Richard A Berk. 2011. "Incarceration, Education and Transition from Delinquency." *Journal of Criminal Justice* 39(4):355–365.
- Butler, Daniel and Charles Crabtree. 2017. "Moving Beyond Measurement: Adapting Audit Stud-

- ies to Test Bias-Reducing Interventions.” *Journal of Experimental Political Science* 4(1):57–67.
- Chetty, Raj, John Friedman, Emmanuel Saez, Nicholas Turner and Danny Yagan. 2017. “Mobility Report Cards: The Role of Colleges in Intergenerational Mobility.”
- URL:** <https://www.nber.org/papers/w23618>
- Coppock, Alexander. 2018. “Avoiding Post-Treatment Bias in Audit Experiments.” *Journal of Experimental Political Science* p. 124.
- Costa, Mia. 2017. “How Responsive are Political Elites? A Meta-Analysis of Experiments on Public Officials*.” *Journal of Experimental Political Science* 4(3):241–254.
- Deming, David and Susan Dynarski. 2010. *College Aid*. University of Chicago Press pp. 283–302.
- Diaz-Serrano, Luis and Enric Meix-Llop. 2016. “Do schools discriminate against homosexual parents? Evidence from a randomized correspondence experiment.” *Economics of Education Review* 53:133 – 142.
- URL:** <http://www.sciencedirect.com/science/article/pii/S0272775716302989>
- Druckman, James N. and Richard M. Shafranek. 2020. “The Intersection of Racial and Partisan Discrimination Evidence from a Correspondence Study of Four-Year Colleges.” *The Journal of Politics* .
- URL:** <https://doi.org/10.1086/708776>
- Dynarski, Susan, C.J. Libassi, Katherine Micheltore and Stephanie Owen. 2018. Closing the Gap: The Effect of a Targeted, Tuition-Free Promise on College Choices of High-Achieving, Low-Income Students. Working Paper 25349 National Bureau of Economic Research.
- URL:** <https://www.nber.org/papers/w25349>
- Einstein, Katherine and David Glick. 2017. “Does Race Affect Access to Government Services? An Experiment Exploring Street-Level Bureaucrats and Access to Public Housing.” *American Journal of Political Science* 61(1):100–116.
- Foucault, M. 1977. *Discipline and Punish: The Birth of the Prison*. Vintage Books.
- Gaddis, Michael. 2014. “Discrimination in the Credential Society: An Audit Study of Race and College Selectivity in the Labor Market.” *Social Forces* 93(4):1451–1479.

- Hemker, Johannes and Anselm Rink. 2017. "Multiple Dimensions of Bureaucratic Discrimination: Evidence From German Welfare Offices." *American Journal of Political Science* 61(4):786–803.
- Hjalmarsson, Randi, Helena Holmlund and Matthew J. Lindquist. 2015. "The Effect of Education on Criminal Convictions and Incarceration: Causal Evidence from Microdata." *The Economic Journal* 125(587):1290–1326.
- Hoxby, Caroline M. 2009. "The Changing Selectivity of American Colleges." *Journal of Economic Perspectives* 23(4):95–118.
- Hoxby, Caroline and Sarah Turner. 2015. "What High-Achieving Low-Income Students Know about College." *American Economic Review* 105(5):514–17.
- Iacus, Stefano, Gary King and Giuseppe Porro. 2012. "Causal Inference Without Balance Checking: Coarsened Exact Matching." *Political Analysis* 20(1):1–24.
- Jilke, Sebastian, Wouter Van Dooren and Sabine Rys. 2018. "Discrimination and Administrative Burden in Public Service Markets: Does a Public–Private Difference Exist?" *Journal of Public Administration Research and Theory* 28(3):423–439.
- Lasswell, H.D. 1936. *Politics: Who Gets What, When, How*. Whittlesey House, McGraw-Hill Book Company, Incorporated.
- Manza, J. and C. Uggen. 2006. *Locked Out: Felon Disenfranchisement and American Democracy*. Studies in Crime and Public Policy Oxford University Press.
- Milkman, Katherine L, Modupe Akinola and Dolly Chugh. 2015. "What Happens Before? A Field Experiment Exploring How Pay and Representation Differentially Shape Bias on the Pathway Into Organizations." *Journal of applied psychology* 100(6):1678–1712.
- Niskanen, William. 1968. "The Peculiar Economics of Bureaucracy." *The American Economic Review* 58(2):293–305.
- Owens, Michael and Adrienne Smith. 2012. "Deviants and Democracy: Punitive Policy Designs and the Social Rights of Felons as Citizens." *American Politics Research* 40(3):531–567.
- Pager, Devah. 2003. "The Mark of a Criminal Record." *American Journal of Sociology* 108(5):937–975.

- Pager, Devah, Bart Bonikowski and Bruce Western. 2009. "Discrimination in a Low-Wage Labor Market: A Field Experiment." *American Sociological Review* 74(5):777–799. PMID: 20689685.
- Pettigrew, Thomas. 1998. "Intergroup Contact Theory." *Annual Review of Psychology* 49(1):65–85.
- Rosenthal, A., E. NaPier, P. Warth and M. Weissman. 2015. Boxed out: Criminal History Screening and College Application Attrition. Technical report Center for Community Alternatives.
URL: <http://communityalternatives.org/pdf/publications/BoxedOutullReport.pdf>
- Stewart, Robert and Christopher Uggen. 2020. "Criminal records and college admissions: A modified experimental audit." *Criminology* 58(1):156–188.
URL: <https://onlinelibrary.wiley.com/doi/abs/10.1111/1745-9125.12229>
- Tversky, Amos and Daniel Kahneman. 1974. "Judgment Under Uncertainty: Heuristics and Biases." *Science* 185(4157):1124–1131.
- Weaver, Vesla and Amy Lerman. 2010. "Political Consequences of the Carceral State." *American Political Science Review* 104(4):817–833.
- White, Ariel, Noah Nathan and Julie Faller. 2015. "What Do I Need to Vote? Bureaucratic Discretion and Discrimination by Local Election Officials." *American Political Science Review* 109(1):129–142.

A Supplementary Information (Online Only)

A.1 Emails

For each of the four names, we created two different emails accounts to reduce the risk of being classified as spam when sending the emails. The email addresses only differ in the order of the three digits that make up the middle part of the address. The inclusion of the three-digit number was necessary to make sure that the email address was not already in use. In Table A3, we list the eight emails that we use to run the experiment. In our estimation of race effects, we pool results across email and name. Individual names and emails did not exhibit measurable effects on our outcomes, conditional on race. This is shown in Table B2.

Table A3: List of Emails

| |
|----------------------------|
| darnell.banks143@gmail.com |
| darnell.banks134@gmail.com |
| tyrone.booker143@gmail.com |
| tyrone.booker134@gmail.com |
| kevin.schmidt143@gmail.com |
| kevin.schmidt134@gmail.com |
| bob.krueger143gmail.com |
| bob.krueger134@gmail.com |

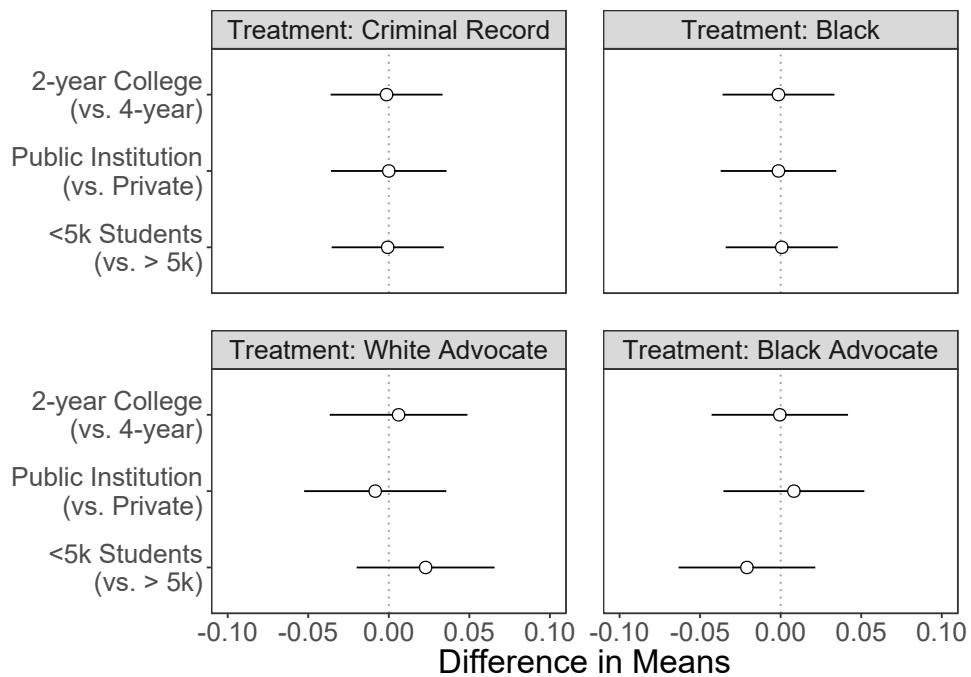
Note: This table shows the list of emails we have created for the experiment. For each applicant name, we create two emails that vary only in the order of the last three digits. We sent an equal proportion of emails from each account.

A.2 CEM & Treatment Assignment

In Section 1, we give a brief overview of the randomization procedure that we employ to guarantee balance across contextual variables. CEM uses pre-treatment variables, coarsens them, and then creates strata based on the coarsened covariate. Since two of the three variables are already binary, CEM can only coarsen the student body size variable. We then use these strata to form pairs of observations that are members of the same stratum. Within pairs, the criminal record, race, and

advocate treatment conditions are randomized so that one unit in the pair receives treatment and one unit receives control for each treatment category. Since the advocate race treatment depends on the presence of an advocate, we do not use pair matching to assign treatment status. Rather, we use simple randomization to determine whether an advocate has a putatively Black or White. We also randomize the name of the email sender or the advocate and the email account from which the message is sent. Conditional on the vector of treatments for each observations, the choice of name and email account is always only between two options. Therefore, we use a simple random draw to determine which email account and which name will be used.

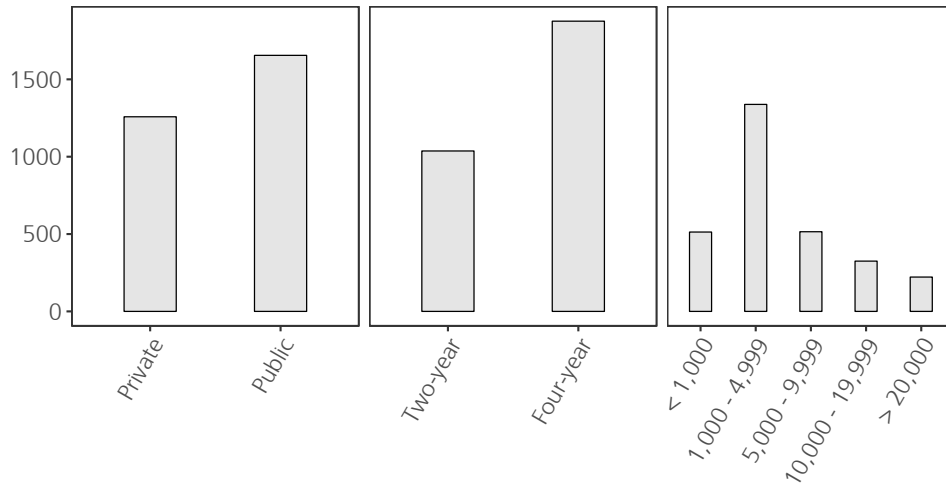
Figure A3: Covariate Balance after CEM



The panels represents the differences in means between treated and untreated units, separately for each covariate. All variables shown in the plot are binary.

B Tables and Figures (Online Only)

Figure B1: School Characteristics



Note: This figure shows the absolute frequencies of different categories of institutions in our sample. The categories in the last panel correspond to the number of students at each institution.

Table B1: Racial Connotation Pre-Test Results

| Applicant name | Correctly classified | CI |
|-----------------------|-----------------------------|----------------|
| Tyrone Booker | 0.95 | [0.924, 0.981] |
| Darnell Banks | 0.90 | [0.862, 0.941] |
| Jamal Gaines | 0.91 | [0.866, 0.944] |
| Kevin Schmidt | 0.95 | [0.918, 0.978] |
| Bob Krueger | 0.96 | [0.936, 0.987] |
| Todd Novak | 0.95 | [0.923, 0.980] |

Note: The table shows how many times each name was correctly classified, i.e. how frequently the names was assigned to the race that we intended it to be perceived as. In total, the names were classified by 200 respondents on MTurk.

Figure B2: Email Instruments

Criminal Record/Advocate Treatment

From: [EMAIL ADDRESS]
To: [ADMISSIONS EMAIL ADDRESS]
Subject: Admissions Info

Hello,

A past student of mine, [APPLICANT NAME], is interested in applying to [SCHOOL], but is worried he is not eligible. He has his GED, which he got at [PENITENTIARY]. Does this affect his eligibility? What else does he need to apply? Are you currently accepting applications?

Thank You,

[INSTRUCTOR NAME]

Criminal Record/ No Advocate

From: [EMAIL ADDRESS]
To: [ADMISSIONS EMAIL ADDRESS]
Subject: Admissions Info

Hello,

I am interested in applying to [SCHOOL], but I am worried I am not eligible. I have my GED, which I got at [PENITENTIARY]. Does this affect my eligibility? What else do I need to apply? Are you currently accepting applications?

Thank You,

[APPLICANT NAME]

No Criminal Record/Advocate Treatment

From: [EMAIL ADDRESS]
To: [ADMISSIONS EMAIL ADDRESS]
Subject: Admissions Info

Hello,

A past student of mine, [APPLICANT NAME], is interested in applying to [SCHOOL], but is worried he is not eligible. He has his GED, which he got online. Does this affect his eligibility? What else does he need to apply? Are you currently accepting applications?

Thank You,

[INSTRUCTOR NAME]

No Criminal Record/ No Advocate

From: [EMAIL ADDRESS]
To: [ADMISSIONS EMAIL ADDRESS]
Subject: Admissions Info

Hello,

I am interested in applying to [SCHOOL], but I am worried I am not eligible. I have my GED, which I got online. Does this affect my eligibility? What else do I need to apply? Are you currently accepting applications?

Thank You,

[APPLICANT NAME]

Note: The figure shows that exact wording of the emails that we sent to the colleges in the sample. In total, there are $2*2*3=16$ different treatment conditions. In this figure, we only show that differences in email wording for the *Criminal Record* and *Advocate* treatments. Applicant and advocate race are signaled using putatively Black or White names, while the email text stays the same.

Table B2: Response Rates and Treatment Distribution by Name

| Name | Response rate | Total emails | % Criminal Record | % Black | % White Advocate | % Black Advocate |
|---------------|---------------|--------------|-------------------|---------|------------------|------------------|
| Bob Krueger | 0.758 | 756 | 50.1 | 0 | 26 | 24.9 |
| Darnell Banks | 0.771 | 721 | 50.2 | 100 | 25.9 | 22.5 |
| Kevin Schmidt | 0.775 | 711 | 51.2 | 0 | 25.3 | 27.6 |
| Tyrone Booker | 0.786 | 746 | 48.5 | 100 | 21.5 | 26.4 |

Note: This tables contains response rates, conditional on the name of the applicant. The last four columns are the relative frequencies of each treatment. For the two advocate treatments, the shares shown in the table signify the probability that an applicant with each respective name was assigned a Black or White advocate. To give an example, 26% of all applicants named ‘Bob Krueger’ had a White advocate, and 24.9% had a Black advocate.

Table B3: Mean Response Rates for the Treatment and Control Groups

| | Response Rate | | Diff. in Means | SE | P-value | $N_{\text{Treatment}}$ | N_{Control} |
|-------------------|----------------|--------------|----------------|-------|---------|------------------------|----------------------|
| | Treatment Mean | Control Mean | | | | | |
| Criminal Record | 0.749 | 0.797 | -0.047 | 0.015 | 0.002 | 1467 | 1467 |
| Black | 0.779 | 0.767 | 0.012 | 0.015 | 0.437 | 1467 | 1467 |
| White Advocate | 0.791 | 0.763 | 0.030 | 0.019 | 0.105 | 724 | 1467 |
| Black Advovcate | 0.776 | 0.763 | 0.012 | 0.019 | 0.522 | 743 | 1467 |
| Advocate (Pooled) | 0.783 | 0.763 | 0.021 | 0.015 | 0.176 | 1467 | 1467 |

Note: This table contains mean response rates conditional on treatment status. The first column indicates the treatment variable. For the advocate treatment, the control mean always refers to the mean response rate that were not sent by an advocate. We also show the standard error of the difference in means between response rates, and the corresponding p-values for the null hypothesis that the true difference is zero.

Table B4: Response Rates for All Possible Treatment Combinations

| Criminal Record Treatment | Race Treatment | Advocate Treatment | Response Rate | N |
|---------------------------|----------------|--------------------|---------------|-----|
| No Criminal Record | White | No Advocate | 0.751 | 360 |
| No Criminal Record | White | White Advocate | 0.856 | 174 |
| No Criminal Record | White | Black Advocate | 0.783 | 190 |
| No Criminal Record | Black | No Advocate | 0.810 | 397 |
| No Criminal Record | Black | White Advocate | 0.808 | 172 |
| No Criminal Record | Black | Black Advocate | 0.803 | 174 |
| Criminal Record | White | No Advocate | 0.748 | 346 |
| Criminal Record | White | White Advocate | 0.756 | 203 |
| Criminal Record | White | Black Advocate | 0.746 | 194 |
| Criminal Record | Black | No Advocate | 0.737 | 364 |
| Criminal Record | Black | White Advocate | 0.747 | 175 |
| Criminal Record | Black | Black Advocate | 0.776 | 185 |

Note: This table reports mean response rates and number of observations for all 12 possible treatment combinations.

Table B5: Main Results – Response Outcome

| | Dependent variable: Response (0/1) | | | | | |
|-------------------|------------------------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| | Full Sample | | | Advocate Only | | |
| Criminal Record | −0.048*** (0.015) | −0.052*** (0.015) | −0.048*** (0.015) | −0.051*** (0.015) | −0.056*** (0.022) | −0.064*** (0.022) |
| Black | 0.012 (0.016) | 0.012 (0.015) | 0.012 (0.016) | 0.012 (0.015) | −0.001 (0.022) | −0.003 (0.022) |
| White Advocate | 0.030 (0.019) | 0.034* (0.019) | | | | |
| Black Advocate | 0.015 (0.019) | 0.011 (0.019) | | | −0.015 (0.022) | −0.021 (0.022) |
| Advocate (Pooled) | | | 0.023 (0.016) | 0.022 (0.015) | | |
| Intercept | 0.779*** (0.016) | 0.769*** (0.111) | 0.779*** (0.016) | 0.768*** (0.111) | 0.820*** (0.022) | 0.856*** (0.143) |
| Covariates | No | Yes | No | Yes | No | Yes |
| State FE | No | Yes | No | Yes | No | Yes |
| N | 2917 | 2917 | 2917 | 2917 | 1459 | 1459 |
| R-squared | 0.004 | 0.041 | 0.004 | 0.040 | 0.005 | 0.062 |

Note: The outcome is a binary response indicator. The treatments are all binary. The covariates are public/private, two-year/four-year and a five-category scale of institution size. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

Table B6: Treatment Effects for Public and Private Schools

| | Dependent variable: Response (0/1) | | |
|--------------------------------------|---|----------------------|----------------------|
| Criminal Record | −0.052*** (0.015) | −0.099*** (0.024) | −0.053*** (0.015) |
| Black | 0.012 (0.015) | 0.009 (0.015) | −0.021 (0.024) |
| Advocate | 0.034* (0.019) | 0.032* (0.019) | 0.033* (0.019) |
| Black Advocate | −0.022 (0.022) | −0.019 (0.022) | −0.022 (0.022) |
| Public Institution (vs. Private) | 0.081*** (0.021) | 0.024 (0.027) | 0.038 (0.027) |
| Criminal Record × Public Institution | | 0.083*** (0.031) | |
| Black × Public Institution | | | 0.057* (0.031) |
| Intercept | 0.769*** (0.111) | 0.723*** (0.110) | 0.705*** (0.110) |
| Covariates | Yes | Yes | Yes |
| State FE | Yes | Yes | Yes |
| N | 2917 | 2917 | 2917 |
| R-squared | 0.041 | 0.045 | 0.044 |

Note: The outcome is a binary response indicator. The treatments are all binary. The covariates are two-year/four-year and a five-category scale of institution size. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

Table B7: Main results with interactions

| | Dependent variable: Response (0/1) | | | |
|----------------------------------|------------------------------------|---------------------|---------------------|----------------------|
| Criminal Record | -0.052*** (0.015) | -0.034 (0.022) | -0.043* (0.022) | -0.052*** (0.015) |
| Black | 0.012 (0.015) | 0.029 (0.022) | 0.012 (0.015) | 0.027 (0.022) |
| White Advocate | 0.034* (0.019) | 0.033* (0.019) | 0.055** (0.027) | 0.061** (0.027) |
| Black Advocate | 0.011 (0.019) | 0.012 (0.019) | 0.009 (0.027) | 0.013 (0.027) |
| Criminal Record × Black | | -0.034 (0.031) | | |
| Criminal Record × White Advocate | | | -0.042 (0.038) | |
| Criminal Record × Black Advocate | | | 0.005 (0.038) | |
| Black × White Advocate | | | | -0.057 (0.038) |
| Black × Black Advocate | | | | -0.003 (0.038) |
| Intercept | 0.769*** (0.111) | 0.761*** (0.112) | 0.760*** (0.112) | 0.757*** (0.112) |
| Covariates | Yes | Yes | Yes | Yes |
| State FE | Yes | Yes | Yes | Yes |
| N | 2917 | 2917 | 2917 | 2917 |
| R-squared | 0.041 | 0.041 | 0.041 | 0.041 |

Note: The outcome is a binary response indicator. The average response rate is 74.4%. The covariates are public/private, two-year/four-year, institution size and state fixed effects. ***p < 0.01; **p < 0.05; *p < 0.1

Table B8: Selected School Characteristics for Private and Public Colleges

| | Mean | |
|-------------------------------------|-------------|-----------|
| | Private | Public |
| Rejection Rate (2013) | 0.378 | 0.336 |
| Avg. SAT (2013) | 1077.414 | 1049.878 |
| Sticker Price in \$ (2013) | 30312.230 | 5780.843 |
| Net Price in \$ (2013) | 17590.030 | 7955.190 |
| Pct. of Students with Parents in Q1 | 0.075 | 0.139 |
| Pct. Black (2000) | 0.109 | 0.132 |
| Parent Median Income in \$ | 96279.550 | 70383.510 |

Note: This tables shows averages for seven school characteristics, separately for public and private colleges. The data is based on Chetty et al. (2017).

Table B9: Effect of a criminal record, conditional on school characteristics

| | Dependent variable: Response (0/1) | | | | | | |
|---|------------------------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| Criminal Record | -0.079*** (0.024) | -0.085*** (0.027) | -0.065*** (0.018) | -0.063*** (0.018) | -0.061*** (0.019) | -0.067*** (0.018) | -0.060*** (0.019) |
| Rejection Rate | -0.007 (0.016) | | | | | | |
| Rejection Rate * Criminal Record | -0.020 (0.024) | | | | | | |
| Avg. SAT | | 0.032* (0.019) | | | | | |
| Avg. SAT * Criminal Record | | 0.022 (0.027) | | | | | |
| Sticker Price | | | 0.006 (0.015) | | | | |
| Sticker Price * Criminal Record | | | -0.019 (0.018) | | | | |
| Net Price | | | | 0.005 (0.015) | | | |
| Net Price * Criminal Record | | | | -0.021 (0.018) | | | |
| Pct. with Parents in Q1 | | | | | -0.054*** (0.014) | | |
| Pct. with Parents in Q1 * Criminal Record | | | | | 0.035* (0.019) | | |
| Pct. Black | | | | | | -0.040*** (0.012) | |
| Pct. Black * Criminal Record | | | | | | -0.016 (0.018) | |
| Parent Median Income | | | | | | | 0.023* (0.013) |
| Parent Median Income * Criminal Record | | | | | | | -0.011 (0.019) |
| Covariates | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Remaining Treatments | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| N | 1,262 | 1,012 | 2,003 | 2,006 | 1,900 | 1,992 | 1,900 |
| R-squared | 0.020 | 0.031 | 0.015 | 0.015 | 0.019 | 0.028 | 0.012 |

Note: The outcome is a binary response indicator. All college characteristics are standardized, such that the coefficients measure one-standard deviation increases. The covariates are two-year/four-year and a five-category scale of institution size. In all models, we also include the remaining two treatments, i.e. the advocate treatment and the applicant race treatment. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

Table B10: Effect of a criminal record, conditional on school characteristics (**Public Schools**)

| Dependent variable: Response (0/1) | | | | | | | |
|---|-------------------|-------------------|--------------------|-------------------|----------------------|---------------------|--------------------|
| Criminal Record | -0.011 (0.037) | -0.022 (0.041) | -0.032 (0.023) | -0.031 (0.023) | -0.025 (0.023) | -0.035 (0.023) | -0.024 (0.023) |
| Rejection Rate | -0.021 (0.026) | | | | | | |
| Rejection Rate * Criminal Record | -0.024 (0.037) | | | | | | |
| Avg. SAT | | 0.044 (0.029) | | | | | |
| Avg. SAT * Criminal Record | | 0.007 (0.041) | | | | | |
| Sticker Price | | | 0.022 (0.018) | | | | |
| Sticker Price * Criminal Record | | | -0.0002 (0.023) | | | | |
| Net Price | | | | 0.005 (0.018) | | | |
| Net Price * Criminal Record | | | | 0.008 (0.023) | | | |
| Pct. with Parents in Q1 | | | | | -0.059*** (0.017) | | |
| Pct. with Parents in Q1 * Criminal Record | | | | | 0.038 (0.023) | | |
| Pct. Black | | | | | | -0.036** (0.015) | |
| Pct. Black * Criminal Record | | | | | | -0.010 (0.023) | |
| Parent Median Income | | | | | | | 0.045** (0.018) |
| Parent Median Income * Criminal Record | | | | | | | -0.015 (0.023) |
| Covariates | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Remaining Treatments | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| N | 452 | 393 | 1,149 | 1,151 | 1,096 | 1,138 | 1,096 |
| R-squared | 0.020 | 0.026 | 0.006 | 0.004 | 0.015 | 0.014 | 0.010 |

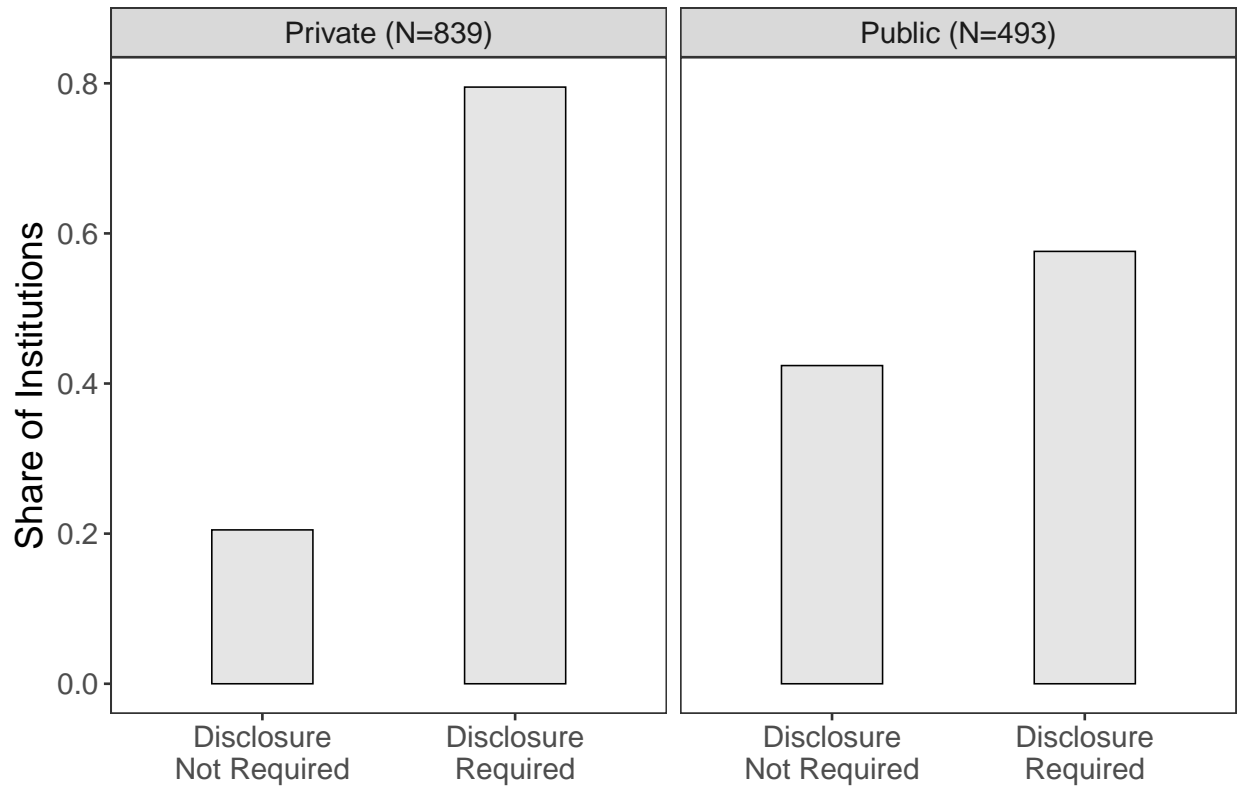
Note: The outcome is a binary response indicator. All college characteristics are standardized, such that the coefficients measure one-standard deviation increases. We limit the sample to public colleges. The covariates are two-year/four-year and a five-category scale of institution size. In all models, we also include the remaining two treatments, i.e. the advocate treatment and the applicant race treatment. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

Table B11: Effect of criminal record conditional on school characteristics (**Private Schools**)

| | Dependent variable: Response (0/1) | | | | | | |
|---|------------------------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| Criminal Record | -0.119*** (0.030) | -0.127*** (0.035) | -0.114*** (0.030) | -0.111*** (0.030) | -0.112*** (0.030) | -0.112*** (0.029) | -0.113*** (0.030) |
| Rejection Rate | -0.005 (0.021) | | | | | | |
| Rejection Rate * Criminal Record | -0.011 (0.031) | | | | | | |
| Avg. SAT | | 0.031 (0.026) | | | | | |
| Avg. SAT * Criminal Record | | 0.036 (0.035) | | | | | |
| Sticker Price | | | -0.005 (0.021) | | | | |
| Sticker Price * Criminal Record | | | 0.054* (0.030) | | | | |
| Net Price | | | | 0.002 (0.020) | | | |
| Net Price * Criminal Record | | | | 0.012 (0.030) | | | |
| Pct. with Parents in Q1 | | | | | -0.040* (0.023) | | |
| Pct. with Parents in Q1 * Criminal Record | | | | | -0.008 (0.030) | | |
| Pct. Black | | | | | | -0.051** (0.021) | |
| Pct. Black * Criminal Record | | | | | | -0.022 (0.029) | |
| Parent Median Income | | | | | | | 0.002 (0.021) |
| Parent Median Income * Criminal Record | | | | | | | 0.033 (0.031) |
| Covariates | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Remaining Treatments | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| N | 810 | 619 | 854 | 855 | 804 | 854 | 804 |
| R ² | 0.025 | 0.043 | 0.032 | 0.025 | 0.033 | 0.047 | 0.026 |

Note: The outcome is a binary response indicator. All college characteristics are standardized, such that the coefficients measure one-standard deviation increases. We limit the sample to **private** colleges. The covariates are two-year/four-year and a five-category scale of institution size. In all models, we also include the remaining two treatments, i.e. the advocate treatment and the applicant race treatment. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

Figure B3: Probability of requiring criminal record disclosure, private and public schools



Note: The figure shows the relative frequency of institutions that have or do not have criminal record disclosure requirements. We only consider schools for which we have data on disclosure requirements, which is about 70% of all four-year colleges.

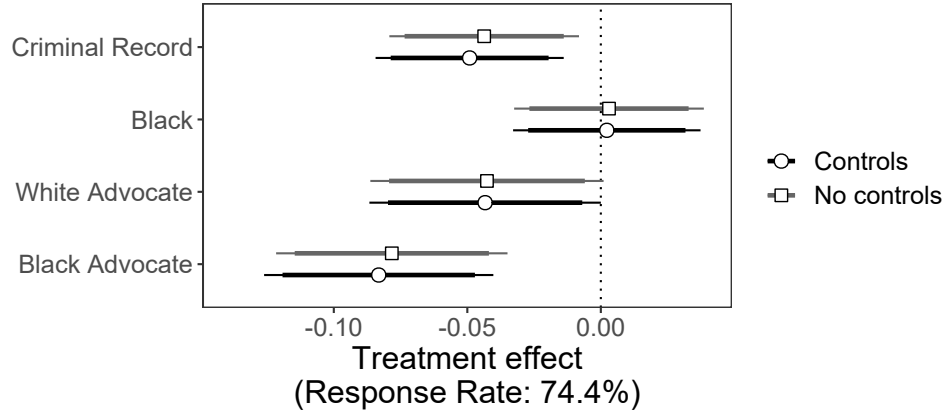
Table B12: Interacting Criminal Record and Disclosure Requirements

| | Dependent Variable: Response (0/1) | | |
|---------------------------------------|---|---------------------|---------------------|
| Criminal Record | -0.022 (0.043) | -0.026 (0.043) | -0.034 (0.044) |
| Disclosure Required | 0.071* (0.037) | 0.069* (0.037) | 0.088** (0.039) |
| Criminal Record * Disclosure Required | -0.088* (0.051) | -0.084 (0.051) | -0.080 (0.052) |
| Intercept | 0.757*** (0.032) | 0.731*** (0.035) | 0.679*** (0.213) |
| Covariates | No | No | Yes |
| State FE | No | No | Yes |
| Remaining Treatments | No | Yes | Yes |
| N | 1330 | 1330 | 1330 |
| R ² | 0.013 | 0.016 | 0.081 |

Note: The outcome is a binary response indicator. The average response rate is 74.4%. The covariates are two-year/four-year and institution size. We only consider schools for which we have data on disclosure requirements, which is about 70% of all four-year colleges. ***p < 0.01; **p < 0.05; *p < 0.1

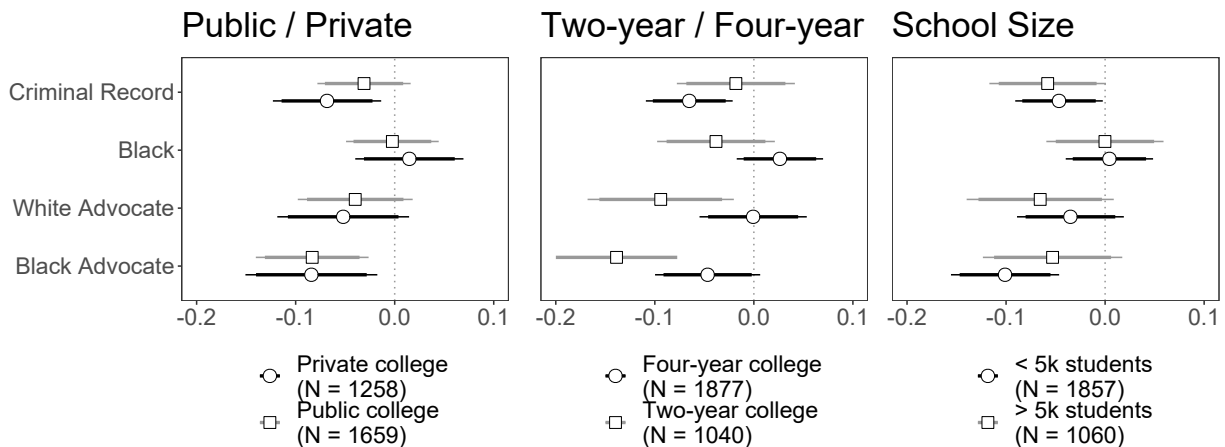
B.1 Friendliness outcome

Figure B4: Main Results – Friendliness



Note: The figure show coefficient estimates from the main specifications. Each pair of coefficients refers to a treatment, which is shown on the y-axis. The outcome is a binary friendliness indicator. Positive effect sizes indicate that the treatment condition increases friendliness. The covariates are public/private, two-year/four-year, institution size and state fixed effects. The solid horizontal lines indicate 90% (thick lines) and 95% (thin lines) confidence intervals.

Figure B5: Results Conditional on School Characteristics – Friendliness



Note: The figures show coefficient estimates by school characteristics. Each pair of coefficients refers to a treatment, which is shown on the y-axis. The outcome is a binary friendliness indicator. Positive effect sizes indicate that the treatment condition increases friendliness. Each panel splits the sample into two groups defined by a school characteristic. We then estimate coefficients separately for the resulting subsamples. All specifications include covariates and state fixed effects. The covariates are public/private, two-year/four-year and institution size. The solid horizontal lines indicate 90% (thick lines) and 95% (thin lines) confidence intervals.

Table B13: Main Results – Friendliness

| Dependent variable: Friendliness (0/1) | | | | | | |
|--|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| | Full sample | | | | Advocate emails only | |
| Criminal Record | −0.044** (0.018) | −0.049*** (0.018) | −0.043** (0.018) | −0.049*** (0.018) | −0.067*** (0.026) | −0.077*** (0.026) |
| Black | 0.003 (0.018) | 0.002 (0.018) | 0.003 (0.018) | 0.002 (0.018) | −0.005 (0.026) | −0.002 (0.026) |
| White Advocate | −0.043* (0.022) | −0.043* (0.022) | | | | |
| Black Advocate | −0.078*** (0.022) | −0.083*** (0.022) | | | −0.036 (0.026) | −0.037 (0.026) |
| Advocate (Pooled) | | | −0.061*** (0.018) | −0.064*** (0.018) | | |
| Intercept | 0.649*** (0.018) | 0.695*** (0.129) | 0.649*** (0.018) | 0.694*** (0.129) | 0.622*** (0.026) | 0.743*** (0.170) |
| Covariates | No | Yes | No | Yes | No | Yes |
| State FE | No | Yes | No | Yes | No | Yes |
| N | 2917 | 2917 | 2917 | 2917 | 1459 | 1459 |
| R-squared | 0.007 | 0.057 | 0.006 | 0.057 | 0.006 | 0.084 |

Note: The outcome is a binary friendliness indicator. The covariates are public/private, two-year/four-year and a five-category scale of institution size. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

Table B14: Main Results with Interactions – Friendliness

| | Dependent variable: Friendliness (0/1) | | | |
|----------------------------------|---|----------------------|---------------------|----------------------|
| Criminal Record | −0.049*** (0.018) | −0.058** (0.025) | −0.025 (0.025) | −0.049*** (0.018) |
| Black | 0.002 (0.018) | −0.007 (0.025) | 0.002 (0.018) | 0.006 (0.025) |
| White Advocate | −0.043* (0.022) | −0.043* (0.022) | −0.019 (0.032) | −0.032 (0.031) |
| Black Advocate | −0.083*** (0.022) | −0.083*** (0.022) | −0.059* (0.031) | −0.087*** (0.031) |
| Criminal Record * Black | | 0.019 (0.036) | | |
| Criminal Record * White Advocate | | | −0.048 (0.044) | |
| Criminal Record * Black Advocate | | | −0.048 (0.044) | |
| Black * White Advocate | | | | −0.024 (0.044) |
| Black * Black Advocate | | | | 0.009 (0.044) |
| Intercept | 0.695*** (0.129) | 0.699*** (0.129) | 0.678*** (0.130) | 0.690*** (0.130) |
| Covariates | Yes | Yes | Yes | Yes |
| State FE | Yes | Yes | Yes | Yes |
| N | 2917 | 2917 | 2917 | 2917 |
| R-squared | 0.057 | 0.058 | 0.058 | 0.058 |

Note: The outcome is a binary friendliness indicator. All regressions include covariates and state fixed effects. The covariates are public/private, two-year/four-year and a five-category scale of institution size. The last model only considers cases where an advocate sent the email. ***p < .01; **p < .05; *p < .1

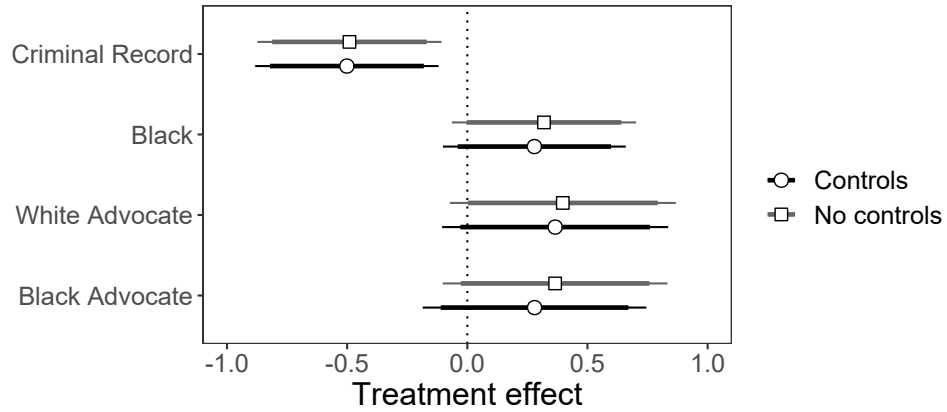
Table B15: Results for Public and Private Schools – Friendliness

| | Dependent variable: Friendliness (0/1) | | |
|--------------------------------------|---|----------------------|----------------------|
| Criminal Record | −0.049*** (0.018) | −0.071*** (0.027) | −0.049*** (0.018) |
| Black | −0.043* (0.022) | −0.045** (0.022) | −0.044** (0.022) |
| White Advocate | −0.083*** (0.022) | −0.082*** (0.022) | −0.082*** (0.022) |
| Black Advocate | 0.002 (0.018) | 0.001 (0.018) | 0.009 (0.027) |
| Public Institution (vs. Private) | 0.035 (0.025) | 0.005 (0.032) | 0.030 (0.031) |
| Criminal Record × Public Institution | | 0.039 (0.036) | |
| Black × Public Institution | | | −0.012 (0.036) |
| Intercept | 0.695*** (0.129) | 0.666*** (0.128) | 0.646*** (0.128) |
| Covariates | Yes | Yes | Yes |
| State FE | Yes | Yes | Yes |
| N | 2917 | 2917 | 2917 |
| R-squared | 0.057 | 0.060 | 0.059 |

Note: The outcome is a binary friendliness indicator. The treatments are all binary. The covariates are public/private, two-year/four-year and a five-category scale of institution size. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

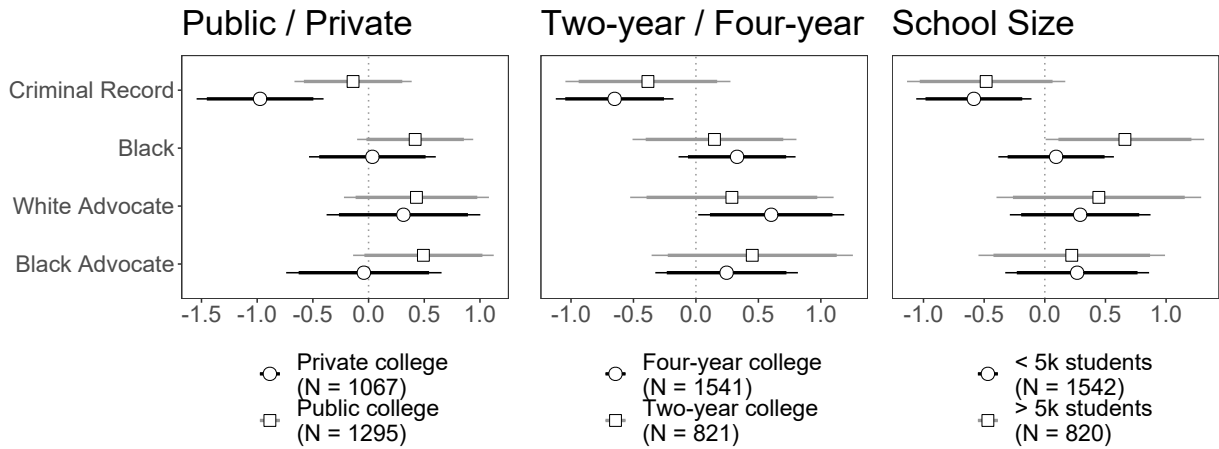
B.2 Thoroughness outcome

Figure B6: Main Results – Thoroughness



Note: The figures show coefficient estimates from the main regressions for the thoroughness outcome. The outcome ranges from 0–3. The covariates are public/private, two-year/four-year and a five-category scale of institution size. The solid horizontal lines indicate 90% (thick lines) and 95% (thin lines) confidence intervals.

Figure B7: Results Conditional on School Characteristics – Thoroughness



Note: The figures show coefficient estimates by school characteristics. Each pair of coefficients refers to a treatment, which is shown on the y-axis. The outcome is thoroughness, which ranges from 0–3. Each panel splits the sample into two groups defined by a school characteristic. We then estimate coefficients separately for the resulting sub-samples. All specifications include covariates and state fixed effects. The covariates are public/private, two-year/four-year and institution size. The solid horizontal lines indicate 90% (thick lines) and 95% (thin lines) confidence intervals.

Table B16: Main Results – Thoroughness

| | Dependent variable: Thoroughness (0–3) | | | | | |
|-------------------|---|---------------------|---------------------|------------------------------|---------------------|---------------------|
| | Full sample | | | Applicant emails only | | |
| Criminal Record | –0.490** (0.196) | –0.501** (0.195) | –0.490** (0.195) | –0.500** (0.195) | –0.578** (0.277) | –0.674** (0.277) |
| Black | 0.319 (0.196) | 0.279 (0.194) | 0.319 (0.195) | 0.278 (0.194) | 0.081 (0.278) | –0.008 (0.275) |
| White Advocate | 0.397* (0.240) | 0.366 (0.240) | | | | |
| Black Advocate | 0.365 (0.239) | 0.280 (0.237) | | | –0.031 (0.277) | –0.102 (0.279) |
| Advocate (Pooled) | | | 0.381* (0.196) | 0.322* (0.195) | | |
| Intercept | 6.141*** (0.196) | 6.803*** (1.360) | 6.141*** (0.196) | 6.803*** (1.360) | 6.697*** (0.279) | 7.662*** (1.760) |
| Covariates | No | Yes | No | Yes | No | Yes |
| State FE | No | Yes | No | Yes | No | Yes |
| N | 2362 | 2362 | 2362 | 2362 | 1187 | 1187 |
| R-squared | 0.005 | 0.055 | 0.005 | 0.055 | 0.004 | 0.088 |

Note: The outcome is a 0–3 thoroughness scale. The covariates are public/private, two-year/four-year and a five-category scale of institution size. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

Table B17: Main Results with Interactions – Thoroughness

| | Dependent variable: Thoroughness (0–3) | | | |
|----------------------------------|---|---------------------|---------------------|----------------------|
| Criminal Record | –0.501** (0.195) | –0.369 (0.276) | –0.395 (0.276) | –0.508*** (0.195) |
| Black | 0.279 (0.194) | 0.409 (0.273) | 0.276 (0.194) | 0.523* (0.276) |
| White Advocate | 0.366 (0.240) | 0.363 (0.240) | 0.544 (0.341) | 0.756** (0.335) |
| Black Advocate | 0.280 (0.237) | 0.283 (0.237) | 0.312 (0.332) | 0.369 (0.335) |
| Criminal Record × Black | | –0.264 (0.391) | | |
| Criminal Record × White Advocate | | | –0.354 (0.479) | |
| Criminal Record × Black Advocate | | | –0.070 (0.475) | |
| Black × White Advocate | | | | –0.805* (0.478) |
| Black × Black Advocate | | | | –0.168 (0.475) |
| Intercept | 6.803*** (1.360) | 6.733*** (1.365) | 6.714*** (1.368) | 6.618*** (1.365) |
| Covariates | Yes | Yes | Yes | Yes |
| State FE | Yes | Yes | Yes | Yes |
| N | 2362 | 2362 | 2362 | 2362 |
| R-squared | 0.055 | 0.055 | 0.055 | 0.056 |

Note: The outcome is a 0–3 thoroughness scale. All regressions include covariates and state fixed effects. The covariates are public/private, two-year/four-year and a five-category scale of institution size. The last model only considers cases where an advocate sent the email.

*** p < .01; ** p < .05; * p < .1

Table B18: Results for Public and Private Schools – Thoroughness

| | Dependent variable: Thoroughness (0–3) | | |
|--------------------------------------|---|----------------------|----------------------|
| Criminal Record | –0.501** (0.195) | –0.936*** (0.290) | –0.505*** (0.195) |
| Black | 0.279 (0.194) | 0.253 (0.194) | 0.087 (0.290) |
| White Advocate | 0.366 (0.240) | 0.338 (0.240) | 0.357 (0.240) |
| Black Advocate | 0.280 (0.237) | 0.283 (0.237) | 0.278 (0.237) |
| Public institution (vs. Private) | 1.175*** (0.270) | 0.649* (0.339) | 0.873** (0.340) |
| Criminal Record * Public Institution | | 0.807** (0.393) | |
| Black * Public Institution | | | 0.348 (0.392) |
| Intercept | 6.803*** (1.360) | 6.396*** (1.346) | 6.162*** (1.341) |
| Covariates | Yes | Yes | |
| State FE | Yes | Yes | |
| N | 2362 | 2362 | 2362 |
| R-squared | 0.055 | 0.060 | 0.058 |

Note: The outcome is a 0–3 thoroughness scale. The covariates are public/private, two-year/four-year and a five-category scale of institution size. Standard errors are shown in parentheses. ***p < .01; **p < .05; *p < .1

B.3 Adjusting for Multiple Comparisons

Given the large number of tests that we conduct, we additionally adjust p-values using the widely-applied method proposed by (Benjamini and Hochberg, 1995). The goal of the method is to control the expected number of incorrectly rejected null hypotheses. To calculate the adjusted p-values, we first collect all hypothesis tests conducted in this paper. This includes all main effects of the three treatments on each of the three outcomes (see e.g. Figure 1), as well interactions between treatments, interactions between school characteristics (see e.g. Tables B7 and B6) as well as treatment effects within sub-samples defined by school characteristics (see e.g. Figure 2). We do not include the estimated coefficients for the school characteristics, since we never directly discuss their effects. In cases where we estimate the same coefficients either using no school controls or conditional on controls, we use both estimates and corresponding p-values. We end up with a total of 142 hypothesis and corresponding p-values. We then adjust p-values using the (Benjamini and Hochberg, 1995) procedure.

We first present adjusted p-values for the two central claims of this paper, i.e. (1) that response rates are lower for formerly incarcerated applicants and (2) that this effect is stronger for private than for public colleges. We present the unadjusted and adjusted p-values in Table B19. We find that the negative effect of criminal records on response rates remains significant at conventional levels after adjusting for multiple comparisons. This is also the case for the effect of criminal records, estimated separately for the subset of private schools. Finally, we find that the adjusted p-value for the interaction between criminal records and public institutions increases to 0.063, slightly greater than the commonly used threshold of $\alpha = 0.05$. In addition, we present adjusted p-values for all remaining hypothesis tests, interactions and subsamples in Table B20.

Table B19: Estimates, P-values and Adjusted P-values for the Main Results

| Term | Sample | Outcome | Controls | Shown in | Estimate | SE | P-val | Adjusted P-val |
|--------------------------------------|----------------------|----------|-------------------|----------|----------|-------|-------|----------------|
| Black | All Colleges | Response | Controls Included | Figure 1 | 0.012 | 0.015 | 0.441 | 0.615 |
| Black * Public Institution | All Colleges | Response | Controls Included | Table B6 | 0.057 | 0.031 | 0.069 | 0.344 |
| Criminal Record | All Colleges | Response | Controls Included | Figure 1 | -0.052 | 0.015 | 0.001 | 0.021 |
| Criminal Record * Public Institution | All Colleges | Response | Controls Included | Table B6 | 0.083 | 0.031 | 0.008 | 0.063 |
| Black | Private Institutions | Response | Controls Included | Figure 2 | -0.020 | 0.025 | 0.420 | 0.615 |
| Criminal Record | Private Institutions | Response | Controls Included | Figure 2 | -0.099 | 0.025 | 0.000 | 0.005 |
| Black | Public Institutions | Response | Controls Included | Figure 2 | 0.034 | 0.019 | 0.079 | 0.344 |
| Criminal Record | Public Institutions | Response | Controls Included | Figure 2 | -0.016 | 0.019 | 0.414 | 0.615 |

Table B20: Estimates, P-values and Adjusted P-values for all results

| Term | Sample | Outcome | Controls | Estimate | SE | P-val | Adjusted P-val |
|----------------------------|--------------|----------|-------------------|----------|-------|-------|----------------|
| Black | All Colleges | Response | Controls Included | 0.012 | 0.015 | 0.441 | 0.615 |
| Black | All Colleges | Response | No Controls | 0.024 | 0.022 | 0.271 | 0.573 |
| Black | All Colleges | Response | No Controls | 0.012 | 0.016 | 0.425 | 0.615 |
| Black * Black Advocate | All Colleges | Response | Controls Included | -0.003 | 0.038 | 0.938 | 0.958 |
| Black * Black Advocate | All Colleges | Response | No Controls | 0.000 | 0.038 | 0.996 | 0.996 |
| Black * Public Institution | All Colleges | Response | Controls Included | 0.057 | 0.031 | 0.069 | 0.344 |
| Black * White Advocate | All Colleges | Response | Controls Included | -0.057 | 0.038 | 0.133 | 0.364 |
| Black * White Advocate | All Colleges | Response | No Controls | -0.052 | 0.038 | 0.173 | 0.437 |
| Black Advocate | All Colleges | Response | Controls Included | 0.011 | 0.019 | 0.547 | 0.681 |
| Black Advocate | All Colleges | Response | No Controls | 0.016 | 0.019 | 0.412 | 0.615 |
| Black Advocate | All Colleges | Response | No Controls | 0.015 | 0.019 | 0.416 | 0.615 |
| Criminal Record | All Colleges | Response | Controls Included | -0.052 | 0.015 | 0.001 | 0.021 |
| Criminal Record | All Colleges | Response | No Controls | -0.048 | 0.015 | 0.002 | 0.024 |
| Criminal Record | All Colleges | Response | No Controls | -0.036 | 0.022 | 0.101 | 0.364 |
| Criminal Record * Black | All Colleges | Response | Controls Included | -0.034 | 0.031 | 0.269 | 0.573 |
| Criminal Record * Black | All Colleges | Response | No Controls | -0.023 | 0.031 | 0.449 | 0.615 |

(continued)

| Term | Sample | Outcome | Controls | Estimate | SE | P-val | Adjusted P-val |
|--------------------------------------|----------------------|----------------|-------------------|-----------------|-----------|--------------|-----------------------|
| Criminal Record * Black Advocate | All Colleges | Response | Controls Included | 0.005 | 0.038 | 0.887 | 0.947 |
| Criminal Record * Black Advocate | All Colleges | Response | No Controls | 0.007 | 0.038 | 0.848 | 0.947 |
| Criminal Record * Public Institution | All Colleges | Response | Controls Included | 0.083 | 0.031 | 0.008 | 0.063 |
| Criminal Record * White Advocate | All Colleges | Response | Controls Included | -0.042 | 0.038 | 0.275 | 0.573 |
| Criminal Record * White Advocate | All Colleges | Response | No Controls | -0.040 | 0.038 | 0.289 | 0.578 |
| White Advocate | All Colleges | Response | Controls Included | 0.034 | 0.019 | 0.079 | 0.344 |
| White Advocate | All Colleges | Response | No Controls | 0.030 | 0.019 | 0.115 | 0.364 |
| White Advocate | All Colleges | Response | No Controls | 0.030 | 0.019 | 0.116 | 0.364 |
| Black | Private Institutions | Response | Controls Included | -0.020 | 0.025 | 0.420 | 0.615 |
| Black Advocate | Private Institutions | Response | Controls Included | 0.006 | 0.031 | 0.844 | 0.947 |
| Criminal Record | Private Institutions | Response | Controls Included | -0.099 | 0.025 | 0.000 | 0.005 |
| White Advocate | Private Institutions | Response | Controls Included | 0.025 | 0.031 | 0.412 | 0.615 |
| Black | Public Institutions | Response | Controls Included | 0.034 | 0.019 | 0.079 | 0.344 |
| Black Advocate | Public Institutions | Response | Controls Included | 0.012 | 0.024 | 0.602 | 0.722 |
| Criminal Record | Public Institutions | Response | Controls Included | -0.016 | 0.019 | 0.414 | 0.615 |
| White Advocate | Public Institutions | Response | Controls Included | 0.043 | 0.024 | 0.071 | 0.344 |
| Black | Two-year Colleges | Response | Controls Included | 0.017 | 0.025 | 0.488 | 0.650 |
| Black Advocate | Two-year Colleges | Response | Controls Included | 0.004 | 0.031 | 0.885 | 0.947 |
| Criminal Record | Two-year Colleges | Response | Controls Included | -0.039 | 0.025 | 0.124 | 0.364 |
| White Advocate | Two-year Colleges | Response | Controls Included | 0.027 | 0.031 | 0.383 | 0.615 |
| Black | Four-year Colleges | Response | Controls Included | 0.012 | 0.020 | 0.554 | 0.681 |
| Black Advocate | Four-year Colleges | Response | Controls Included | 0.015 | 0.024 | 0.543 | 0.681 |
| Criminal Record | Four-year Colleges | Response | Controls Included | -0.062 | 0.020 | 0.002 | 0.024 |
| White Advocate | Four-year Colleges | Response | Controls Included | 0.047 | 0.024 | 0.053 | 0.344 |
| Black | > 5k Students | Response | Controls Included | 0.028 | 0.024 | 0.250 | 0.573 |
| Black Advocate | > 5k Students | Response | Controls Included | -0.003 | 0.029 | 0.924 | 0.958 |
| Criminal Record | > 5k Students | Response | Controls Included | -0.038 | 0.024 | 0.122 | 0.364 |
| White Advocate | > 5k Students | Response | Controls Included | 0.030 | 0.031 | 0.333 | 0.615 |
| Black | < 5k Students | Response | Controls Included | 0.005 | 0.020 | 0.822 | 0.947 |
| Black Advocate | < 5k Students | Response | Controls Included | 0.019 | 0.025 | 0.440 | 0.615 |
| Criminal Record | < 5k Students | Response | Controls Included | -0.061 | 0.020 | 0.002 | 0.024 |
| White Advocate | < 5k Students | Response | Controls Included | 0.037 | 0.025 | 0.137 | 0.364 |

(continued)

| Term | Sample | Outcome | Controls | Estimate | SE | P-val | Adjusted P-val |
|--------------------------------------|----------------------|----------------|-------------------|-----------------|-----------|--------------|-----------------------|
| Black | All Colleges | Thoroughness | Controls Included | 0.279 | 0.194 | 0.150 | 0.344 |
| Black | All Colleges | Thoroughness | No Controls | 0.319 | 0.196 | 0.103 | 0.323 |
| Black | All Colleges | Thoroughness | No Controls | 0.364 | 0.274 | 0.184 | 0.385 |
| Black * Black Advocate | All Colleges | Thoroughness | Controls Included | -0.168 | 0.475 | 0.724 | 0.821 |
| Black * Black Advocate | All Colleges | Thoroughness | No Controls | -0.161 | 0.477 | 0.735 | 0.821 |
| Black * Public Institution | All Colleges | Thoroughness | Controls Included | 0.348 | 0.392 | 0.375 | 0.567 |
| Black * White Advocate | All Colleges | Thoroughness | Controls Included | -0.805 | 0.478 | 0.093 | 0.323 |
| Black * White Advocate | All Colleges | Thoroughness | No Controls | -0.793 | 0.480 | 0.098 | 0.323 |
| Black Advocate | All Colleges | Thoroughness | Controls Included | 0.280 | 0.237 | 0.239 | 0.458 |
| Black Advocate | All Colleges | Thoroughness | No Controls | 0.366 | 0.239 | 0.125 | 0.323 |
| Black Advocate | All Colleges | Thoroughness | No Controls | 0.365 | 0.239 | 0.126 | 0.323 |
| Criminal Record | All Colleges | Thoroughness | Controls Included | -0.501 | 0.195 | 0.010 | 0.146 |
| Criminal Record | All Colleges | Thoroughness | No Controls | -0.490 | 0.196 | 0.012 | 0.146 |
| Criminal Record | All Colleges | Thoroughness | No Controls | -0.445 | 0.277 | 0.108 | 0.323 |
| Criminal Record * Black | All Colleges | Thoroughness | Controls Included | -0.264 | 0.391 | 0.500 | 0.667 |
| Criminal Record * Black | All Colleges | Thoroughness | No Controls | -0.091 | 0.391 | 0.816 | 0.890 |
| Criminal Record * Black Advocate | All Colleges | Thoroughness | Controls Included | -0.070 | 0.475 | 0.883 | 0.925 |
| Criminal Record * Black Advocate | All Colleges | Thoroughness | No Controls | -0.037 | 0.477 | 0.939 | 0.939 |
| Criminal Record * Public Institution | All Colleges | Thoroughness | Controls Included | 0.807 | 0.393 | 0.040 | 0.287 |
| Criminal Record * White Advocate | All Colleges | Thoroughness | Controls Included | -0.354 | 0.479 | 0.460 | 0.650 |
| Criminal Record * White Advocate | All Colleges | Thoroughness | No Controls | -0.298 | 0.480 | 0.535 | 0.694 |
| White Advocate | All Colleges | Thoroughness | Controls Included | 0.366 | 0.240 | 0.128 | 0.323 |
| White Advocate | All Colleges | Thoroughness | No Controls | 0.397 | 0.240 | 0.098 | 0.323 |
| White Advocate | All Colleges | Thoroughness | No Controls | 0.396 | 0.240 | 0.099 | 0.323 |
| Black | Private Institutions | Thoroughness | Controls Included | 0.034 | 0.290 | 0.906 | 0.925 |
| Black Advocate | Private Institutions | Thoroughness | Controls Included | -0.042 | 0.356 | 0.905 | 0.925 |
| Criminal Record | Private Institutions | Thoroughness | Controls Included | -0.974 | 0.291 | 0.001 | 0.040 |
| White Advocate | Private Institutions | Thoroughness | Controls Included | 0.313 | 0.352 | 0.374 | 0.567 |
| Black | Public Institutions | Thoroughness | Controls Included | 0.418 | 0.266 | 0.116 | 0.323 |
| Black Advocate | Public Institutions | Thoroughness | Controls Included | 0.492 | 0.322 | 0.127 | 0.323 |
| Criminal Record | Public Institutions | Thoroughness | Controls Included | -0.139 | 0.268 | 0.604 | 0.744 |
| White Advocate | Public Institutions | Thoroughness | Controls Included | 0.430 | 0.332 | 0.196 | 0.391 |

(continued)

| Term | Sample | Outcome | Controls | Estimate | SE | P-val | Adjusted P-val |
|----------------------------------|--------------------|----------------|-------------------|-----------------|-----------|--------------|-----------------------|
| Black | Two-year Colleges | Thoroughness | Controls Included | 0.149 | 0.335 | 0.657 | 0.789 |
| Black Advocate | Two-year Colleges | Thoroughness | Controls Included | 0.451 | 0.411 | 0.273 | 0.485 |
| Criminal Record | Two-year Colleges | Thoroughness | Controls Included | -0.384 | 0.337 | 0.255 | 0.470 |
| White Advocate | Two-year Colleges | Thoroughness | Controls Included | 0.288 | 0.415 | 0.488 | 0.667 |
| Black | Four-year Colleges | Thoroughness | Controls Included | 0.329 | 0.239 | 0.168 | 0.367 |
| Black Advocate | Four-year Colleges | Thoroughness | Controls Included | 0.246 | 0.291 | 0.399 | 0.581 |
| Criminal Record | Four-year Colleges | Thoroughness | Controls Included | -0.651 | 0.240 | 0.007 | 0.146 |
| White Advocate | Four-year Colleges | Thoroughness | Controls Included | 0.603 | 0.298 | 0.043 | 0.287 |
| Black | > 5k Students | Thoroughness | Controls Included | 0.660 | 0.333 | 0.048 | 0.287 |
| Black Advocate | > 5k Students | Thoroughness | Controls Included | 0.222 | 0.392 | 0.571 | 0.722 |
| Criminal Record | > 5k Students | Thoroughness | Controls Included | -0.483 | 0.333 | 0.147 | 0.344 |
| White Advocate | > 5k Students | Thoroughness | Controls Included | 0.445 | 0.430 | 0.301 | 0.516 |
| Black | < 5k Students | Thoroughness | Controls Included | 0.093 | 0.243 | 0.701 | 0.821 |
| Black Advocate | < 5k Students | Thoroughness | Controls Included | 0.267 | 0.303 | 0.378 | 0.567 |
| Criminal Record | < 5k Students | Thoroughness | Controls Included | -0.585 | 0.242 | 0.016 | 0.153 |
| White Advocate | < 5k Students | Thoroughness | Controls Included | 0.292 | 0.296 | 0.324 | 0.537 |
| Black | All Colleges | Friendliness | Controls Included | 0.002 | 0.018 | 0.900 | 0.975 |
| Black | All Colleges | Friendliness | No Controls | -0.007 | 0.026 | 0.775 | 0.938 |
| Black | All Colleges | Friendliness | No Controls | 0.003 | 0.018 | 0.866 | 0.972 |
| Black * Black Advocate | All Colleges | Friendliness | Controls Included | 0.009 | 0.044 | 0.836 | 0.972 |
| Black * Black Advocate | All Colleges | Friendliness | No Controls | -0.004 | 0.044 | 0.933 | 0.975 |
| Black * White Advocate | All Colleges | Friendliness | Controls Included | -0.024 | 0.044 | 0.584 | 0.749 |
| Black * White Advocate | All Colleges | Friendliness | No Controls | -0.026 | 0.045 | 0.563 | 0.749 |
| Black Advocate | All Colleges | Friendliness | Controls Included | -0.083 | 0.022 | 0.000 | 0.004 |
| Black Advocate | All Colleges | Friendliness | No Controls | -0.078 | 0.022 | 0.000 | 0.004 |
| Black Advocate | All Colleges | Friendliness | No Controls | -0.078 | 0.022 | 0.000 | 0.004 |
| Criminal Record | All Colleges | Friendliness | Controls Included | -0.049 | 0.018 | 0.006 | 0.036 |
| Criminal Record | All Colleges | Friendliness | No Controls | -0.044 | 0.018 | 0.016 | 0.061 |
| Criminal Record | All Colleges | Friendliness | No Controls | -0.054 | 0.026 | 0.035 | 0.123 |
| Criminal Record * Black | All Colleges | Friendliness | Controls Included | 0.019 | 0.036 | 0.602 | 0.749 |
| Criminal Record * Black | All Colleges | Friendliness | No Controls | 0.021 | 0.036 | 0.566 | 0.749 |
| Criminal Record * Black Advocate | All Colleges | Friendliness | Controls Included | -0.048 | 0.044 | 0.269 | 0.430 |

(continued)

| Term | Sample | Outcome | Controls | Estimate | SE | P-val | Adjusted P-val |
|----------------------------------|----------------------|----------------|-------------------|-----------------|-----------|--------------|-----------------------|
| Criminal Record * Black Advocate | All Colleges | Friendliness | No Controls | -0.044 | 0.044 | 0.323 | 0.479 |
| Criminal Record * White Advocate | All Colleges | Friendliness | Controls Included | -0.048 | 0.044 | 0.280 | 0.430 |
| Criminal Record * White Advocate | All Colleges | Friendliness | No Controls | -0.049 | 0.045 | 0.275 | 0.430 |
| White Advocate | All Colleges | Friendliness | Controls Included | -0.043 | 0.022 | 0.050 | 0.144 |
| White Advocate | All Colleges | Friendliness | No Controls | -0.043 | 0.022 | 0.056 | 0.144 |
| White Advocate | All Colleges | Friendliness | No Controls | -0.043 | 0.022 | 0.056 | 0.144 |
| Black | Private Institutions | Friendliness | Controls Included | 0.015 | 0.028 | 0.595 | 0.749 |
| Black Advocate | Private Institutions | Friendliness | Controls Included | -0.084 | 0.034 | 0.013 | 0.061 |
| Criminal Record | Private Institutions | Friendliness | Controls Included | -0.068 | 0.028 | 0.014 | 0.061 |
| White Advocate | Private Institutions | Friendliness | Controls Included | -0.052 | 0.034 | 0.126 | 0.276 |
| Black | Public Institutions | Friendliness | Controls Included | -0.002 | 0.024 | 0.922 | 0.975 |
| Black Advocate | Public Institutions | Friendliness | Controls Included | -0.083 | 0.029 | 0.004 | 0.028 |
| Criminal Record | Public Institutions | Friendliness | Controls Included | -0.031 | 0.024 | 0.198 | 0.367 |
| White Advocate | Public Institutions | Friendliness | Controls Included | -0.040 | 0.030 | 0.178 | 0.357 |
| Black | Two-year Colleges | Friendliness | Controls Included | -0.038 | 0.030 | 0.208 | 0.367 |
| Black Advocate | Two-year Colleges | Friendliness | Controls Included | -0.139 | 0.037 | 0.000 | 0.004 |
| Criminal Record | Two-year Colleges | Friendliness | Controls Included | -0.018 | 0.030 | 0.550 | 0.749 |
| White Advocate | Two-year Colleges | Friendliness | Controls Included | -0.094 | 0.038 | 0.013 | 0.061 |
| Black | Four-year Colleges | Friendliness | Controls Included | 0.026 | 0.022 | 0.237 | 0.404 |
| Black Advocate | Four-year Colleges | Friendliness | Controls Included | -0.047 | 0.027 | 0.085 | 0.195 |
| Criminal Record | Four-year Colleges | Friendliness | Controls Included | -0.065 | 0.022 | 0.004 | 0.027 |
| White Advocate | Four-year Colleges | Friendliness | Controls Included | -0.001 | 0.028 | 0.979 | 0.996 |
| Black | > 5k Students | Friendliness | Controls Included | 0.000 | 0.030 | 0.996 | 0.996 |
| Black Advocate | > 5k Students | Friendliness | Controls Included | -0.053 | 0.036 | 0.140 | 0.292 |
| Criminal Record | > 5k Students | Friendliness | Controls Included | -0.058 | 0.030 | 0.054 | 0.144 |
| White Advocate | > 5k Students | Friendliness | Controls Included | -0.066 | 0.038 | 0.084 | 0.195 |
| Black | < 5k Students | Friendliness | Controls Included | 0.004 | 0.022 | 0.847 | 0.972 |
| Black Advocate | < 5k Students | Friendliness | Controls Included | -0.101 | 0.028 | 0.000 | 0.004 |
| Criminal Record | < 5k Students | Friendliness | Controls Included | -0.047 | 0.023 | 0.039 | 0.129 |
| White Advocate | < 5k Students | Friendliness | Controls Included | -0.035 | 0.028 | 0.203 | 0.367 |